

High rate of adaptive evolution in two widespread European pines

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Comparing related organisms with differing ecological requirements and evolutionary 36 37 histories can shed light on the mechanisms and drivers underlying genetic adaptation. Here, by examining a common set of hundreds of loci, we compare patterns of nucleotide 38 diversity and molecular adaptation of two European conifers (Scots pine and maritime pine) 39 living in contrasted environments and characterized by distinct population genetic structure 40 (low and clinal in Scots pine, high and ecotypic in maritime pine) and demographic 41 histories. We found higher nucleotide diversity in Scots pine than in maritime pine, whereas 42 rates of new adaptive substitutions (ω_a), as estimated from the Distribution of Fitness 43 Effects (DFE), were similar across species, and among the highest found in plants. Sample 44 size and population genetic structure did not appear to have resulted in any significant bias 45 46 in ω_a . Moreover, the species-specific population contraction-expansion dynamics did not 47 seem to have affected differentially the rate of adaptive substitution in these two pines. 48 Several methodological and biological factors may underlie the unusually high rate of adaptive evolution of Scots pine and maritime pine. By providing two new case studies with 49 contrasting evolutionary history, we contribute to disentangling the multiple factors 50 potentially affecting adaptive evolution in natural plant populations. 51

52 Introduction

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Understanding of the mechanisms of plant adaptation has been advanced through 54 55 comparative studies of species differing in demographic history (Slotte et al. 2010), effective population size (Gossmann et al. 2012; Strasburg et al. 2011), genetic structure 56 among populations (Wright & Andolfatto 2008), ecological conditions (Tellier et al. 2011), 57 or phylogenetic relationships (Eckert et al. 2013a; Grivet et al. 2013; Palmé et al. 2009). 58 For long-lived species such as most forest trees, unraveling adaptive processes is 59 challenging, as during their long life-cycle, individuals experience different selective 60 61 pressures that accumulate over time. Genomic research in forest trees, moreover, has been hindered because of large genome sizes and the lack of model species. Nevertheless, recent 62 advances in tree genomics and reanalysis of common garden experiments have fostered a 63 64 body of literature that provides insights into the mechanisms underlying forest tree adaptation in space and time. Emerging from these studies are the ideas that forest trees are 65 66 locally adapted to different environmental conditions, especially to temperature, photoperiod, drought, or biotic stress (see examples in Alberto et al. 2013a; Savolainen et 67 al. 2007), and that they respond to contrasted selection pressure across life stages (Alía et 68 al. 2014). Based on this knowledge, it is of prime interest to identify major environmental 69 70 drivers of adaptation, as well as the genes involved in the process, as they can help forecast 71 the future distribution of these ecologically important species in the face of climate change.

Identifying footprints of natural selection within genomes is complex, as the observed patterns of polymorphism may result from many distinct yet interacting evolutionary forces, including neutral processes such as migration or genetic drift, all acting on variation generated by mutation. Because demography and selection can leave similar patterns within genomes (Biswas & Akey 2006; Excoffier *et al.* 2009; Nielsen 2005), tests

aiming at attributing these patterns to the action of natural selection must take into account 77 78 the demographic history of the species, as well as the complex interactions between these two processes (Li et al. 2012; Schrider & Kern 2016). Tests robust to demography often 79 contrast polymorphism within species with divergence between species. For example, the 80 popular McDonald-Kreitman (MK) test (McDonald & Kreitman 1991) compares the 81 82 amount of polymorphism to divergence for categories of sites that are expected to evolve differently (e.g., synonymous/silent vs. non-synonymous). From the original MK test, 83 several extensions have been developed. Two of the most popular are the MKPRF test 84 (Bustamante et al. 2002, 2003, 2005), which is more powerful but is based on specification 85 86 of a population genetic model, and the more recent SnIPRE test (Eilertson et al. 2012), which is a nonparametric method that does not require estimation of population genetic 87 88 parameters.

89 Other methods based on the MK test provide overall estimates of adaptive evolution that can be correlated with population parameters (e.g., effective population size; Gossmann 90 91 et al. 2012) or environmental characteristics (e.g., Lourenço et al. 2013) in comparative 92 studies. First, the Distribution of Fitness Effects (DFE) of new mutations at functional sites is estimated from polymorphism data and under a specific demographic scenario. Second, 93 the number of substitutions originating from neutral and slightly deleterious mutations is 94 95 predicted from the estimated DFE (e.g., the method II of Eyre-Walker & Keightley 2009). Any excess of substitutions (with respect to the neutral expectation) can then be attributed 96 to directional selection, as measured by α (the proportion of adaptive nucleotide 97 substitutions) or ω_a (the relative rate of adaptive substitutions scaled by the rate of neutral 98 substitution) (Bierne & Eyre-Walker 2004; Eyre-Walker & Keightley 2009; Smith & Eyre-99 100 Walker 2002). Although α has been extensively used (e.g., (Eckert *et al.* 2013a; Gossmann et al. 2010; Slotte et al. 2010; Strasburg et al. 2011), this estimate is of limited value for 101

estimating the efficiency of the adaptive process, because α is also influenced by the rate of non-adaptive substitutions (Gossmann *et al.* 2012). The parameter ω_a , which is roughly equivalent to K_a (the number of non-synonymous substitutions per non-synonymous site), is thus more appropriate for comparing adaptive evolution across genomic regions or species (Gossmann *et al.* 2012; Lourenço *et al.* 2013).

In the present study, we assess the effect of distinct demographic histories and 107 selection regimes on inferences of adaptive evolution for two widespread conifer species. 108 109 Scots pine (*Pinus sylvestris* L.) is widely and continuously distributed in Eurasia, occupying regions that differ greatly in climate (Krakau et al. 2013). Its demographic history is 110 111 characterized by an ancient bottleneck (Kujala & Savolainen 2012; Pyhäjärvi et al. 2007) and limited population genetic structure, which is only found along the margins of its wide 112 range (Gullberg 1985; Kujala & Savolainen 2012; Pyhäjärvi et al. 2007). The absence of 113 114 genetic structure shown by molecular markers across much of its range, however, is not indicative of a lack of quantitative trait variation across populations (Andersson & 115 116 Fedorkov 2004; García-Gil et al. 2003; Notivol et al. 2007; Oleksyn et al. 1998). In 117 particular, strong clines in photoperiod-related traits are often observed across latitudinal gradients in the species (e.g., for timing of growth cessation and budset; see Beck et al. 118 2004; Hurme et al. 1997; Mikola 1982; Oleksyn et al. 1998). 119

In contrast to Scots pine, maritime pine (*Pinus pinaster* Ait.) is patchily distributed across the western Mediterranean Basin and the Atlantic regions of Portugal, Spain and France. This conifer grows in warm temperate regions with an oceanic influence on climate (Abad Viñas *et al.* 2016), and is particularly well adapted to dry and fire-prone environments. Its demographic history is characterized by a more recent bottleneck relative to Scots pine (Naydenov *et al.* 2014), and it has a strong genetic structure among populations across its range (e.g., Burban and Petit 2003; Bucci et al. 2007; Jaramillo-

Correa et al. 2015) that is accompanied by morphological and physiological differences
(Kremer and Roussel 1986; Alía et al. 1995; Santos del Blanco et al. 2010; Lamy et al.
2011; Corcuera et al. 2012), defining various subspecies and eco-types (Richardson 1998).
This genetic structure likely results from post-Pleistocene events (Burban and Petit 2003;
Bucci et al. 2007; Naydenov et al. 2014), including adaptation to local climate (SerraVarela et al. 2015).

Given the large differences between Scots pine and maritime pine with respect to 133 134 their climatic niches, we focused on genes for general responses to different type of abiotic stress. This is also consistent with previous studies ranking the importance of different 135 136 fitness-related traits for local adaptation in the species (for Scots pine see Castro et al. 2002; Galiano et al. 2010; Ryyppö et al. 1998; and for maritime pine see Corcuera et al. 2011; 137 Gaspar et al. 2013; Lamy et al. 2014; Ramírez-Valiente & Robledo-Arnuncio 2014). Then, 138 139 we used complementary methods to get insights into the action of selective forces, both at 140 specific genes and genome-wide, and considered the specific demographic, ecological and 141 historical settings of each species to discuss the possible factors (both methodological and 142 biological) that may explain our results. Our hypotheses are: i) Scots pine has high nucleotide diversity resulting from its large distribution and expected large effective 143 population size; ii) Large effective population size in Scots pine would have also resulted in 144 145 higher efficiency of selection and thus a higher number of fixed adaptive substitutions (see 146 also Gossmann et al. 2012); and iii) the highly fragmented distribution in maritime pine would have resulted in lower levels of overall adaptive evolution. Scots pine and maritime 147 148 pine, with their contrasting characteristics, allow for exploring how different biological 149 factors may interact with natural selection and adaptive evolution in plants.

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152 Materials and methods

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154 Sampling

155 Scots pine was sampled across 8-10 populations from its western range (Figure 1 and Table S1) for eight loci from six candidate genes for abiotic stress response and phenology (3-20 156 individuals per population, and a total of 36-115 individuals per locus), and across a smaller 157 number of populations (7) for the much larger CRSP (Comparative Re-Sequencing in 158 159 Pinaceae initiative; http://dendrome.ucdavis.edu/NealeLab/crsp/; see Wegrzyn et al. 2008) set of 364 loci (1-2 individuals per population, and a total of 6-12 individuals per locus). In 160 161 a first phase, maritime pine was sampled across 10-12 populations from its full range (Figure 1 and Table S1) for the same eight abiotic stress response and phenology loci as for 162 Scots pine (1-19 individuals per population for a total of 49-100 individuals per locus), and 163 164 across 11 populations for the CRSP gene dataset (1-3 individuals per population for a total of 6-14 individuals per locus). More details on the gene datasets are presented in the 165 166 Sequence data section. In a second phase, another dataset produced in maritime pine 167 comprising 128 loci common with the CRSP dataset (2-4 individuals per population for a total of 28-36 individuals per locus) was obtained within the Conifer Re-sequencing 168 Initiative for European Conifers (CRIEC, www.evoltree.eu; Figure 1 and Table S1). This 169 170 dataset with more individuals per population was examined in order to study the effect of 171 sample size and population structure on the DFE-based estimates. Finally, the CRSP and the CRIEC datasets were combined for the 128 loci in common (2-6 individuals per population 172 173 for a total of 34-49 individuals per locus) to reach a bigger sample size for the statistical 174 analyses ("extended CRSP" dataset). Sequence datasets used for the different analyses are 175 detailed in Figure S1.

Trees were sampled within populations following standard protocols to avoid sampling related trees (i.e., leaving a minimum distance of 50 m between sampled individuals) and without any phenotypic selection.

179

180 *Sequence data*

181 <u>Sequence alignment and editing</u>

Eight loci from previously studied candidate genes, including two full-length genes (coL1 182 183 and gia), were amplified with available primers from different sources (see Table S2). Another 1,600 gene amplicons (hereafter referred as the CRSP gene set) were obtained by 184 resequencing in seven conifers (including Scots pine and maritime pine) some loci 185 originally developed in loblolly pine, Pinus taeda (Eckert et al. 2013b). For both gene sets, 186 DNA sequences were obtained by direct sequencing from haploid seed megagametophytes. 187 In this way, (i) phase is directly known and does not need to be estimated and (ii) co-188 189 amplification of paralogs, a common problem in plant species with large genomes such as 190 conifers, is more easily detected. For the CRSP gene set, loci with at least six successfully 191 sequenced individuals in both Scots pine and maritime pine were accepted for further analyses. This led to 491 loci that were subsequently checked manually and edited with 192 Codon Code Aligner (CodonCode Corporation, Centerville, MA, USA) or Sequencher 4.7 193 194 (Gene Codes Corporation, Ann Arbor, MI, USA). Low quality sequences were removed, as 195 well as putative paralogous genes based on phylogenies including the three pines (the two 196 target species together with loblolly pine), as detected by PRANK runs with default 197 parameters (Löytynoja & Goldman 2005). This last quality-filtering step led to a set of 389 loci common to the two target species. 198

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200 <u>Annotation</u>

201 Gene annotation for all loci was obtained from homology with loblolly pine EST contigs 202 and the NCBI reference protein database using Geneious version 6.1 (Biomatters, Auckland, New Zealand). All 389 genes had high sequence similarity with loblolly pine 203 204 genes, suggesting low copy-number genes and orthology across the different species. The biological function of the genes was determined based on their homology with the 205 Arabidopsis thaliana and/or Pinus taeda protein database with an E-value threshold of 10⁻¹⁰ 206 **S**3) 207 (see Table and the Protein Knowledgebase UniProtKB 208 (http://www.uniprot.org/uniprot/). Of the 389 genes, 364 could be annotated, leading to a total of 372 annotated loci in the two species when adding the eight previously-studied 209 candidate gene loci. Loblolly pine was used as outgroup when needed. 210

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212 <u>Target and reference genes</u>

The 372 loci under study, as well as their annotations, are publicly available (NCBI 213 GenBank: MF385275-MF385581 and MF385585-MF397901). From these 372 loci, 64 214 215 were selected as target genes, including the eight previously-studied candidate gene loci 216 (see Table S2 for details and references) and 56 other annotated loci selected from the CRSP dataset (see above). All 64 target genes had highly confident annotation associated to 217 adaptive traits: 48 loci related to abiotic stress responses (mostly cold, heat, drought, salt 218 219 and other oxidative stresses); and 16 loci related to photosynthesis and photosystem (9 loci), 220 and phenology (7 loci). Only target genes were tested for footprints of selection.

The <u>reference genes</u> consisted of all other loci (308) from the CRSP dataset, including loci with unknown function or loci for which there is no evidence of involvement in forest tree adaptation. Because of implementation limitations and statistical power requirements (i.e., the minimum number of required segregating sites) in the MKPRF test (see below), reference genes were combined for this analysis into compound loci. All loci

with a known position on a genetic map (loblolly pine, Eckert et al. 2010; maritime pine, 226 227 Chancerel et al. 2011; and our own unpublished genetic map of Scots pine) were grouped according to the linkage groups forming 28 compound loci (see Table S4). To balance the 228 number of loci used in the MKPRF test (64 target versus 28 compound reference loci), we 229 used also 36 other reference loci selected at random among those reference genes not 230 included in the compound loci to also reach 64 loci for the reference group in which the 231 statistical power was mainly driven by the compound loci (higher number of segregating 232 233 sites).

234

235 Nucleotide diversity, genetic divergence and overall patterns of polymorphism

Nucleotide diversity π (Nei & Li 1979) and Watterson's θ_w (Watterson 1975) were calculated for all sites, as well as separately for silent, synonymous and non-synonymous sites, in Scots pine and maritime pine. For each species, the divergence from loblolly pine was characterized by the number of synonymous substitutions per synonymous site (K_s) and of non-synonymous substitutions per non-synonymous site (K_a), both with Jukes-Cantor correction, and their ratio (K_a/K_s), as well as the statistics of shared and fixed segregating sites requested by some of the neutrality tests (see below).

To identify overall differences in patterns of polymorphism across species, different 243 244 statistics were computed based on the observed Site Frequency Spectrum (SFS): Tajima's D 245 (Tajima 1989), which evaluates the difference between low- and intermediate-frequency 246 variants; Zeng et al.'s E (Zeng et al. 2006), which evaluates the difference between lowand high-frequency variants; and the normalized Fay and Wu's $H(H_n; \text{Zeng et al. 2006})$, 247 248 which evaluates the difference between intermediate- and high-frequency variants. These 249 statistics were computed, separately, for the 64 target loci and the 308 reference loci, and compared with those under a standard neutral equilibrium model (constant population size). 250

All nucleotide diversity statistics were computed using MANVa and mstatspop (https://bioinformatics.cragenomica.es/numgenomics/people/sebas/software/software.html), and, when needed, the SDMTools package in R was used to correct for sample size variation across loci, by weighting means and standard deviations.

255

256 Neutrality tests robust to demography

To identify potential genes under selection, we first used the SnIPRE approach (Eilertson et 257 258 al. 2012), which considers polymorphism and divergence data from synonymous/silent and non-synonymous sites under a Poisson Random Effect model. This method is based on the 259 McDonald-Kreitman test, and it can reliably identify genes under weak and strong negative 260 261 as well as positive selection, without requiring the specification of a population genetic 262 model. We considered both the selection (specific selection effect of a gene relative to 263 neutrality) and constraint (proportion of non-synonymous mutations that are non-lethal, thus 264 having effects on counts) effects provided by the program. The selection effect is useful to 265 identify selection on mildly deleterious and advantageous mutations, while the constraint 266 effect is useful to identify strong negative or purifying selection (Eilertson et al. 2012). Thus, significant effects on sequence data are classified as being neutral, negative or 267 positive. 268

Second, we used the McDonald-Kreitman Poisson Random Fields (MKPRF) test (Bustamante *et al.* 2002, 2005), which implements a Markov chain Monte Carlo (MCMC) Bayesian approach to estimate parameters of the Poisson Random Field (PRF) model (Sawyer & Hartl 1992), allowing to compare groups of genes for selection signatures. MKPRF not only identifies non-neutrally evolving loci but it also estimates the associated strength of selection ($\gamma=2N_es$). The MKPRF analysis was run using the *mkprf* program (kindly provided by Carlos D. Bustamante and Adam Boyko), with the following

parameters: 10 independent MCMC chains per run and 10,000 samples from each chain
drawn on every 10 steps, after burn-in of 1,000 steps. Default values were used for prior
distributions and other MCMC parameters. Convergence of the MCMC runs was checked
in the MCMC output files generated by the program.

280

281 Distribution of fitness effects (DFE) and adaptive evolution

We used a method based on the DFE to estimate the overall rates of adaptive substitutions (ω_a): the method II of Eyre-Walker and Keightley (2009) as implemented in the DoFE 3.1 program. This method accounts for demographic changes that can affect the shape of the SFS by comparing the observed folded SFS at neutral sites to the folded SFS expected from neutral mutations in a stationary population at equilibrium and assumes that demography has the same proportional effect on the SFS of selected sites.

Input files were created using Python and Perl scripts kindly provided by Jaakko 288 289 Tyrmi and Adrian Schneider, respectively (available upon request). To build the input files 290 for the method II of Eyre-Walker and Keightley (2009), we computed the statistics for non-291 synonymous sites and synonymous sites, at the intraspecific level (for Scots pine and maritime pine, respectively) and at the interspecific level (using loblolly pine as outgroup 292 for each target pine) using the same number of alleles per species and dataset: 305 loci with 293 294 10 alleles each for Scots pine and 291 loci with 11 alleles for maritime pine for the CRSP 295 datasets; 126 loci with 23 alleles each for the CRIEC dataset. The other input statistics were 296 then computed with MANVa (see above) using the same reduced datasets.

Loblolly pine was used as outgroup for both Scots pine and maritime pine and, thus, DFE-based estimates of adaptive evolution for these two pine species were done along a partially shared branch. In phylogenetically close species, shared divergence with the outgroup may result in differences in ω_a mainly due to differences in within-species

polymorphism, i.e., estimates can be biased if the nucleotide divergence between the 301 species is low relative to within species variation (Keightley & Eyre-Walker 2012). To 302 exclude this potential source of bias, shared divergence between loblolly pine and Scots 303 304 pine/maritime pine was estimated using two approaches. First, one, three or five sequences from each locus and species were randomly sampled and concatenated, leading to three 305 aligned fasta files. Aligned fasta files were subsequently used to build unrooted 306 phylogenetic trees using PhyML (Guindon & Gascuel 2003). The analyses were run with 307 308 default parameters, using the General Time Reversible (GTR) model (Tavaré 1986), and for all sites, and both 0-fold and 4-fold degenerate sites. The proportion of shared divergence 309 310 was estimated then as the fraction of the sum of branch lengths due to the common branch leading to loblolly pine (i.e., the common branch leading to loblolly pine divided by the 311 312 total sum of branch lengths). Second, fixed differences between loblolly pine and both Scots pine and maritime pine were counted, and shared fixed differences were directly obtained 313 314 for all sites as well as for 0-fold and 4-fold degenerate sites.

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316

317 **Results**

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319 Nucleotide diversity, genetic divergence and overall patterns of polymorphism

The two pines showed a significant difference (non-overlapping 95% confidence intervals) in nucleotide diversity, with Scots pine ($\theta_{w-syn}=0.00687$, 95% CI: 0.00594-0.00781) displaying 1.64 times the nucleotide diversity of maritime pine ($\theta_{w-syn}=0.00419$, 95% CI: 0.00362-0.00476) (Table 1). In both species, the overall K_a , K_s , and K_a/K_s ratios (relative to loblolly pine) were similar, with overlapping 95% confidence intervals (data not shown), indicating equal divergence from loblolly pine (K_s of 0.0323 for Scots pine and of 0.0375 for maritime pine) and similar rates of evolutionary constraint (Table 1). Scots pine and maritime pine only shared about 2-4% of their synonymous polymorphic sites (depending on the species used as outgroup).

329 Statistics based on the SFS (Fay & Wu 2000; Tajima 1989; Zeng et al. 2006) revealed different polymorphism patterns in the two pines (Table 2). Scots pine displayed 330 an excess of low-frequency variants compared to intermediate-frequency (Tajima's D) and 331 high-frequency (Zeng et al.'s E) variants, but no differences between intermediate- and 332 333 high-frequency (Fay and Wu's H_n) variants. In contrast, maritime pine displayed no difference of intermediate- with respect to low-frequency variants (Tajima's D), and an 334 excess of high-frequency variants compared to low-frequency (Zeng et al.'s E) and 335 intermediate-frequency (Fay and Wu's H_n) variants. Thus, overall, Scots pine displayed an 336 337 excess of low-frequency variants, while maritime pine showed an excess of high-frequency 338 variants. This pattern is present in both target and reference loci suggesting that 339 demographic processes (and not selective processes) underlie the observed differences 340 between the two pines. The impact of population structure on the statistics is reflected 341 especially by Fay and Wu's H_n , which is more sensitive to population subdivision (Zeng et al. 2006), and that deviates more strongly from the standard neutral model in the species 342 with stronger population structure (i.e., maritime pine). 343

344

345 *Neutrality tests robust to demography*

The robust-to-demography SnIPRE method (Empirical Bayes and Bayesian SnIPRE tests) suggested four and three target genes as being under selection in Scots pine (3 positive and 1 negative selection events) and maritime pine (2 positive and 1 negative selection events), respectively (Table 3 and Table S5). Overall, the 64 target loci had a positive average selection effect in Scots pine, while in maritime pine such effect was not observed (Figure

S2). All genes found under selection with the MKPRF test but one for Scots pine (five and 351 352 two for Scots pine and maritime pine, respectively) were among those revealed by the SnIPRE method (Table 3). As expected, a higher proportion of loci from the previously-353 354 studied candidate gene dataset (see Material and Methods) was found under selection (50% and 25% based on the two methods in Scots pine and maritime pine, respectively) compared 355 to the CRSP dataset (0.27% and 0.55%, respectively). Two of these genes were common to 356 both pines using the two methods, with the same pattern of negative (*dhn1*) and positive 357 (coL1) selection. 358

359

360 Distribution of fitness effects (DFE) and adaptive evolution

The inferred DFE using DoFE was similar for the two pines, with most of the mutations 361 being strongly deleterious and subject to purifying selection (Figure 2). However, there 362 363 were also some differences in the DFE spectrum. Scots pine had a significantly higher proportion of deleterious mutations for the three classes with the least deleterious mutations, 364 365 while the opposite pattern was observed at the other extreme of the DFE spectrum with about 77% of mutation being highly deleterious for Scots pine and about 85% for maritime 366 pine. Increasing the sample size for maritime pine (the "extended CRSP" dataset, see 367 Sampling section) resulted in higher proportion of the three classes with the least deleterious 368 mutations but still in lower proportion compared to Scots pine, while the pattern was the 369 370 opposite for highly deleterious mutations (Figure 2). In terms of adaptive substitutions, both species displayed a relative rate significantly different from zero (Scots pine $\omega_a = 0.1156$ 371 vs. maritime pine $\omega_a = 0.1535$), with no significant differences between them (overlapping 372 95% CIs, see Figure 2). Increasing the sampling size for maritime pine ("extended CRSP" 373 374 dataset) led to a slightly smaller ω_a (0.1447), not different from that estimated with the CRSP datasets for Scots pine and maritime pine (Figure 2). Estimating ω_a with a simple 375

and non-parametric method $\omega_a = \frac{d_N - d_S(\frac{P_N}{P_S})}{d_S}$ (Gossmann *et al.* 2010; Kousathanas *et al.* 2014) gave $\omega_a = -0.0740$ with 95% CIs (-0.2430; 0.0871) for Scots pine and $\omega_a = 0.0976$ with 95% CIs (-0.0778; 0.2867) for maritime pine. These estimates confirmed the trend of a higher ω_a in maritime pine than in Scots pine, albeit differences were still not significant (i.e., overlapping CIs).

Shared divergence between loblolly pine and Scots pine/maritime pine was 381 382 estimated using two approaches. First, a three-species unrooted phylogenetic tree was built using 168,534 bp of concatenated sequences, which showed only 43% of shared divergence 383 due to the common branch from loblolly pine to Scots pine/maritime pine (Figure S3). A 384 similar estimate was obtained for 0-fold and 4-fold degenerate sites (43 to 44% of shared 385 386 divergence). Furthermore, we obtained the same estimates with one, three or five sequences randomly-sampled from each locus and species. Second, counts of fixed differences 387 between loblolly pine and Scots pine/maritime pine showed 629 (all sites), 152 (0-fold 388 degenerate sites) and 106 (4-fold degenerate sites) mutations specific to the Scots pine 389 390 lineage, while these numbers were 960, 196 and 185 mutations, respectively, for maritime 391 pine (Table S6).

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394 Discussion

395

The comparison of the same sets of loci between Scots pine and maritime pine, two related species that occupy different ecological niches and are characterized by different evolutionary histories, allows for exploring how different biological factors may interact with natural selection and adaptive evolution in plants. We discuss how our results are in agreement with each species evolutionary history, as well as how the comparison between

401 species helps to better understand the evolutionary forces responsible for the observed402 molecular pattern.

403

404 *Nucleotide diversity*

This study presents the first large set of loci sequenced in a range-wide sample of two 405 important European conifers. The primers used in this study were transferred from a related 406 New World species, loblolly pine, and thus may suffer from ascertainment bias due to 407 408 enrichment with low-diversity conserved genes. To test this hypothesis, we compared a set of 364 loblolly pine loci orthologous to those used in this study with 1,000 sets of 364 loci 409 410 randomly selected with replacement from a larger set of ca. 6,000 loci available in this species (see Eckert et al. 2013b). This comparison indicated a significant bias towards 411 lower nucleotide diversity in our gene set (i.e., non-overlapping 95% confidence intervals), 412 but only when considering statistics for all sites (see Table S7). This suggests that 413 414 comparisons based on synonymous/silent sites or even non-synonymous sites in our study 415 are subject to only limited bias. Moreover, because the ascertainment bias would affect both 416 species in the same way, it would not prevent a comparative analysis between the two pines, which is the main focus of this study. 417

The sequencing of the same 372 loci in both pines revealed significantly higher 418 419 nucleotide diversity in the widespread and continuously distributed Scots pine than in the 420 narrower and more patchily distributed maritime pine. Only a limited number of genes has been sequenced for each pine (and rarely the same genes across species) in previous studies, 421 422 making interspecific comparisons difficult. Watterson's nucleotide diversity for silent sites 423 (compare with Table S8) in Scots pine was 0.00525 for 16 loci (Pyhäjärvi et al. 2007), and 424 0.00620 for 11 loci (Kujala & Savolainen 2012); while that of maritime pine was 0.00824 for 11 loci (Eveno et al. 2008), but was only 0.00280 for six other loci (Grivet et al. 2011) 425

To test whether using a smaller gene set could lead to biased nucleotide diversity estimates, 426 we subsampled different number of loci from the total 372 in our study, and compared 427 nucleotide diversity for these subsamples. Subsampling for 10, 50 or 100 loci did not affect 428 429 average nucleotide diversity estimates (see confidence intervals in Table S9), suggesting that the number of genes used did not necessarily lead to bias. Nevertheless, subsets of 430 specific genes may indeed have substantially different levels of nucleotide diversity. For 431 example, the set of six previously-studied candidates genes (see *Material and methods*) 432 433 showed higher nucleotide diversity (Scots pine $\theta_{w-silent}=0.00895$; maritime pine θ_{w} silent=0.00664) than the 372-locus average (Table S8). These results show that nucleotide 434 diversity can only be properly compared across species when using a large set of 435 (preferably) common genes, as this estimate can be very variable across small specific gene 436 437 sets.

438

439 Insights into demographic history

440 Patterns within the SFS for each species (as summarized by statistics such as Tajima's D, 441 Zeng et al.'s E, and Fay and Wu's H_n) can provide insights into demographic history (see references in Gravel et al. 2011). In Scots pine and maritime pine, they reflected relatively 442 well what is known for each pine species. Scots pine has likely been through a very ancient 443 444 (~1-2 Ma) and severe (shrinking populations to about 1% of present time population size) 445 bottleneck (Pyhäjärvi et al. 2007). In more recent times, the species would have recolonized 446 a vast territory in northern, central and Eastern Europe, reaching northern Fennoscandia some 10,000 to 7,000 years ago (see Pyhäjärvi et al. 2007, 2008; Savolainen et al. 2011; 447 Kujala et al. 2012; Cheddadi et al. 2006; Naydenov et al. 2007). This wide-range 448 449 colonization process would have led to a weak population genetic structure in Scots pine (except especially at the southern margins) (Cheddadi et al. 2006; Dvornyk et al. 2002; 450

García-Gil et al. 2003; Karhu et al. 1996; Pyhäjärvi et al. 2007). The excess of low-451 frequency variants in this species could have originated during repeated long-range 452 expansions and would correspond (mostly) to relatively new mutations (but likely before 453 454 the most recent ice age). In contrast, maritime pine has likely survived in multiple glacial refugia (Bucci et al. 2007; Burban & Petit 2003), from which it would have recolonized its 455 current range around the Last Glacial Maximum (~20,000 years ago; Naydenov et al. 2014). 456 Its spatially limited expansion, combined with population fragmentation, would have led to 457 458 distinct and regionally restricted gene pools (Bucci et al. 2007; Burban & Petit 2003; Jaramillo-Correa et al. 2015), which would have been relatively stable across time, leading 459 460 to the accumulation of high-frequency variants when considering full-range SFS patterns. Because of the very distinct demographic histories of the two pines, we used methods 461 integrating demography to characterize their pattern of adaptive evolution at the molecular 462 463 level and to test loci for signatures of positive and negative selection.

464

465 *Genes under selection*

466 Three dehydrins were found to be under selection in the two pines. Previous studies looking at various members of this family, and using different approaches (neutrality tests, F_{ST} -467 outlier detection, allele frequency-environment correlations), also showed the action of 468 natural selection on dehydrin genes: *dhn1*, *dhn3* and *dhn9* in Scots pine (Palmé *et al.* 2009; 469 Wachowiak et al. 2009); dhn1 (Eveno et al. 2008), dhn2 and dhn5 (Grivet et al. 2011) in 470 maritime pine. In this study, two new dehydrin genes (dhn2 and dhn5) were identified as 471 472 possible targets of positive selection in Scots pine, while our results confirmed the adaptive role of *dhn1* in maritime pine. Gene expression of dehydrins in water stress experiments 473 474 pointed to their involvement in drought resistance in maritime pine (Perdiguero et al. 2012; Velasco-Conde et al. 2012). They have also been shown to be involved in wounding, cold, 475

and drought stress response in white spruce and loblolly pine (Lorenz *et al.* 2011; Richard *et al.* 2000; Watkinson *et al.* 2003). Altogether these studies point to the pivotal role of dehydrins in conifer adaptive response to abiotic stress, over the short- (F_{ST}) and long-term (SnIPRE) timescale, emphasizing their importance.

An interesting case is that of locus 0_4042_01 which showed a signal of positive 480 selection in maritime pine (16 fixed non-synonymous mutations compared to only three in 481 Scots pine). Locus $0_{4042}01$ shows high similarity (E-value of 2×10^{-21}) with glutathione 482 S-transferase (gst) in Arabidopsis thaliana, an enzyme involved in secondary metabolism 483 response to the processes of detoxification and stress response to cold (Goulas et al. 2006), 484 485 salt (Jiang et al. 2007), and pathogens (Jones et al. 2006). In the Chinese pine, Pinus tabuliformis, five residues within gst were found under positive selection, four of them 486 involved in the enzyme activity and specificity (Lan et al. 2013). Interestingly, one of these 487 488 residues codes for a different amino acid in maritime pine (a proline) compared to loblolly pine (alanine), while in Scots pine most individuals maintain the putative ancestral form. 489 490 This amino acid (positions 31-33) is located very close to the catalytically active G-site in 491 the spatial conformation of the protein, and the substitution of this residue could cause structural changes in the GSH binding pocket (where the conjugation of intracellular 492 glutathione to a wide variety of molecules occurs; Lan et al. 2013). 493

Finally, *constans-like 1* (*coL1*), which was found under positive selection in both pines, codes for a putative transcription factor suggested to affect flower development in *Arabidopsis* (Ledger *et al.* 2001). In trees, homologues to *constans* are involved in bud development (Alberto *et al.* 2013b; Ruttink *et al.* 2007), as well as in photoperiodic control of shoot elongation (Holefors *et al.* 2009). Evidence of selection on *constans*-like genes, moreover, has been found in poplar (Chen *et al.* 2014; Ma *et al.* 2010; Ruttink *et al.* 2007;

Smith *et al.* 2004; Wei *et al.* 2013), spruce (Holliday *et al.* 2010), oak (Lind-Riehl *et al.*2014), and the perennial *Arabidopsis lyrata* (Mattila *et al.* 2016).

502

503 Distribution of fitness effects (DFE) and adaptive evolution

Scots pine and maritime pine case studies provided new insights on conifer adaptive 504 evolution and the evolutionary forces shaping conifer genomes. The DFE was similar 505 between the two pines and to other plants (e.g., Eckert et al. 2013a; Gossmann et al. 2010), 506 507 as well as to various other organisms (see references in Eyre-Walker & Keightley 2007), with most of the mutations being strongly deleterious and subject to purifying selection. 508 509 Compared to other conifers (and plants) however (see Eckert et al. 2013a, and Hodgins et al. 2016), both targeted pines presented an atypical pattern with a lower proportion of 510 slightly deleterious mutations and a higher proportion of highly deleterious mutations. The 511 512 relative rates of new adaptive substitutions (ω_a) for Scots pine (0.1156) and maritime pine 513 (0.1535) were within the range found in other species (between -0.14 and 0.31; Gossmann 514 et al. 2012), although at the upper range limit for plants (Gossmann et al. 2010, 2012) as 515 well as higher by a factor over two than those estimated with the same methodology in other pines (ω_a = from -0.0477 to 0.0325 for 11 species of soft pines in Eckert *et al.* 2013a; 516 $\omega_a = 0.0592$ for lodgepole pine in Hodgins *et al.* 2016). 517

Several outcomes emerge from these results. First, sampling intensity may bias the estimate of DFE and ω_a , as illustrated by the "extended CRSP dataset" that led to weaker differences between the two pines. Second, albeit very distinct in terms of their evolutionary histories (i.e., demographic history, population structure, and effective population size), the two pines present similar rates of adaptive evolution. This suggests that other factors may also govern the efficacy of selection across these taxa (see below). Third, it is noteworthy to highlight the high efficiency of natural selection at purging highly deleterious mutations as well as the high rate of positive selection in both pines, in regards to plants in general and more specifically to other conifers studied so far. Plants tend to have low rates of adaptive evolution, linked to contracting populations and a high level of population structure (Gossmann *et al.* 2010). Since pines present in general large effective population size, being also relatively undomesticated species, it is expected that they display also high rates of adaptive evolution. However, this trend was not found in previous studies including pines, suggesting that other factors, still largely unknown, may also be relevant (Chen *et al.* 2017).

532 The reliability of our estimates may depend on methodological factors, among them the choice of the loci under study, the outgroup species, or the analytical method, as well as 533 534 the sampling intensity and distribution (Eckert et al. 2013a; Phifer-Rixey et al. 2012; Städler et al. 2009). More specifically, (i) because we used mostly primers transferred from 535 another species (loblolly pine), gene sets in our study may be more conserved and thus 536 537 undergo less adaptive evolution (Bachtrog 2008; Eckert et al. 2013a; Gossmann et al. 2010); notice that this would make even more remarkable the high ω_a found in both pines. 538 539 (ii) Different genic regions may be under different evolutionary constraints (e.g., Hodgins et 540 al. 2016), and therefore lead to different estimates of adaptive evolution. We computed DFE and ω_a for silent sites (intron + untranslated regions, UTR) and found lower rates of 541 positive selection (data not shown) although not significantly different from those estimated 542 543 with synonymous sites. (iii) Recombination and mutation rates may differ among selected 544 loci and thus can directly affect the proportion of segregating sites (Bachtrog 2008). (iv) The level of divergence of the species of interest with the outgroup may reveal different 545 546 proportion of segregating sites, i.e., the closer the outgroup the fewer differences will be detected (Gossmann et al. 2010; Strasburg et al. 2011). In our study, low phylogenetic 547 548 distance and partial sharing by the targeted species of the branch conducting to the outgroup could have overestimated ω_a (Keightley & Eyre-Walker 2012). However, phylogenetic 549

analyses indicated less than 50% of shared divergence between loblolly pine and each of the 550 551 targeted species suggesting that observed patterns were not only due to within-species polymorphism. Furthermore, counts of fixed differences between the outgroup and each of 552 553 the targeted species pointed to sufficient level of lineage specific mutations in both Scots pine and maritime pine as to correctly estimate ω_a . (v) The choice of the methodology used 554 to infer adaptive evolution (Eckert et al. 2013a; Welch 2006) may give different output. By 555 selecting the same loci and outgroup, as well as the same methodology for the two pines 556 557 and the main studies in other organisms, we attempted to control for these factors and ensured that the estimates (although probably conservative, see point (i)) were comparable 558 across species. Finally, both sampling intensity and distribution may have affected our 559 estimate of adaptive evolution rate. (vi) To test effects of sampling intensity, in terms of 560 both individuals and loci, first, the CRSP dataset (less samples per population but more 561 562 loci) was compared with the "extended CRSP" dataset (more samples per population but 563 fewer loci) in maritime pine and, second, simulated datasets with different number of loci 564 were compared. Neither approach suggested any effect of sampling intensity on our 565 estimates. (vii) Sampling scheme influences what aspects of the history are emphasized (see Städler et al. 2009). Scots pine populations were sampled from both the main range 566 (Finland, Sweden and Poland) and the margins (Spain, Italy and UK), representing fairly 567 568 well the species evolutionary history overall. Nevertheless, there may still be a bias in Scots pine estimates, as the CRSP loci were not sampled across the full species range (in 569 particular the eastern-northern range was poorly sampled). Thus, some apparently fixed 570 571 non-synonymous mutations may show polymorphism in the unsampled range, resulting in 572 an upward bias. In maritime pine, however, this bias should be minimal (if any) as 573 populations were sampled across its full range, considering all gene pools known in the species. In addition, population genetic structure does not seem to have affected the 574

proportion of adaptive substitutions, with estimates being similar between the weakly 575 structured Scots pine and the highly structured maritime pine, and still higher than 576 published estimates for other species. Finally, other factors may constrain selective forces, 577 578 as suggested by comparative studies and theoretical work, and may have influenced the estimates of adaptive evolution in the two pines (Galtier 2016; Lanfear et al. 2014): genetic 579 draft and background selection (Messer & Petrov 2013; Peischl et al. 2013), environmental 580 heterogeneity (Tellier *et al.* 2011), phenotypic space dimensionality (i.e., fitness landscape) 581 582 and rate of environmental change (Gillespie 2001; Lourenço et al. 2013). These factors are challenging to tease apart. 583

584

585 *Conclusion*

By analyzing a common set of 372 gene loci, we detected specific patterns of molecular 586 587 evolution and adaptation in two widespread European conifers. First, as expected, 588 nucleotide diversity was higher in the continuously distributed Scots pine than in the 589 patchily distributed maritime pine. Second, by using methods that incorporate demographic 590 effects, we detected an unexpected high relative rate of adaptive substitutions in both pines, and in particular in maritime pine. Although we cannot fully discard methodological 591 592 caveats, these high rates of adaptive evolution do not seem to be correlated with population 593 genetic structure nor demographic histories that differ between the two pines. Altogether, 594 our results suggest that more than one factor may be responsible of the high rate of adaptive evolution found in these two emblematic pine species, with several factors being difficult to 595 596 tease apart.

597

598

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MF385585-MF397901.

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960 Author Contribution

961
962 DG, OS and SCGM conceived the study. AJE, DBN and SCGM designed and produced
963 the sequence data sets. DG and KA analyzed the data and drafted the manuscript. AV
964 and SCGM contributed to analyze the data. All the authors contributed to editing and

965 revising the manuscript.

966 Tables

967

Table 1. Summary statistics for 372 common loci sequenced in Scots pine and maritime pine.

969

	Scots pine	maritime pine
Number of loci	372	372
Total length (bp)	165,048	166,641
Segregating sites	1543	983
Average length (bp)	444	448
$\theta_{w-\text{syn}}$ (stdev)	0.0069 (0.0090)	0.0042 (0.0054)
$\pi_{\rm syn}$ (stdev)	0.0063 (0.0098)	0.0044 (0.0069)
θ_{w-nsyn} (stdev)	0.0015 (0.0027)	0.0008 (0.0014)
π_{nsyn} (stdev)	0.0013 (0.0025)	0.0009 (0.0018)
K_s (stdev)	0.0323 (0.0253)	0.0375 (0.0313)
K_a (stdev)	0.0070 (0.0081)	0.0070 (0.0080)
K_a / K_s (stdev)	0.3988 (1.0366)	0.3335 (0.8254)

970

971 bp=base pair

972 stdev=standard deviation

973 syn=synonymous

974 nsyn=nonsynonymous

975 θ_w : average Watterson's nucleotide diversity (Watterson 1975) per site

976 π : average Tajima's nucleotide diversity (Tajima 1989) per site

977 K_s : number of synonymous substitutions per synonymous site with Jukes-Cantor correction, using *P. taeda* as outgroup

978 K_a : number of non-synonymous substitutions per non-synonymous site with Jukes-Cantor correction, using *P. taeda* as outgroup

- Table 2. SFS-based statistics for 64 target and 308 reference loci in Scots pine and maritime pine, with their 95% Confidence Interval
 in square brackets.

002								
902	Locus set Statistic		Scots pine	maritime pine				
	Target	Tajima's D	-0.5230 [-0.8559,-0.1901]	0.1195 [-0.2020,0.4410]				
	Target	Zeng et al.'s E	-0.3596 [-0.6483,-0.0709]	0.5309 [0.1473,0.9145]				
	Target	Fay and Wu's H_n	-0.0844 [-0.2788,0.1100]	-0.5269 [-0.9309,-0.1229]				
	Reference	Tajima's D	-0.4776 [-0.5831,-0.3720]	0.0106 [-0.1122,0.1335]				
	Reference	Zeng et al.'s E	-0.3376 [-0.4517,-0.2235]	0.2898 [0.1566,0.4230]				
	Reference	Fay and Wu's H_n	0.0610 [-0.0483,0.1703]	-0.3077 [-0.4587,-0.1567]				

Table 3. Neutrality tests and type of selection for target genes in Scots pine and maritime pine. Symbols in parenthesis represent genes related to biotic stress (S) and

phenology/photosystem (P).

Locus	Scots	pine	maritin	maritime pine			
	SnIPRE*	MKPRF	SnIPRE*	MKPRF			
dhn1 (S)	negative	negative		negative			
<i>dhn2</i> (S)	positive	positive					
<i>dhn5</i> (S)	positive	positive					
coL1 (P)	positive	positive	positive				
0_4042_01 (S)			positive	positive			
2_9480_01 (S)			negative				
0_12156_02 (P)		positive					

*Only genes in common between the Empirical Bayes and the Bayesian SnIPRE tests are

reported. Outputs for each method are presented in full in Table S5.

- 991 Figure Legends
- 992

Figure1. Species distribution and sampling for Scots pine (blue, black dots) and maritimepine (green, red dots).

995

996 Figure 2. Distribution of Fitness Effects of new mutations (DFE) (a), and relative rate of adaptive substitutions ω_a (b) for the CRSP dataset in Scots pine and maritime pine, and for 997 the "extended CRSP" dataset (only maritime pine), using the method II of Eyre-Walker and 998 Keightley (2009), as implemented in DoFE. N_es denotes the product of the effective 999 population size N_e and the strength of selection s, with $N_e s < 1$ corresponding to slightly 1000 deleterious mutations and N_{es} >100 corresponding to highly deleterious mutations. Bars in 1001 (a) represent standard errors (with different letters indicating significant differences), while 1002 bars in (b) represent 95% Confidence Intervals. 1003

Figure1.







- 1 Figure S1. Molecular datasets used for the different analyses. Species abbreviations are as
- 2 follows: Loblolly pine (Pita), Scots pine (Pisy) and maritime pine (Pipn).



Figure S2. Selection events detected with the SnIPRE approach for the 64 target loci and silent sites in Scots pine and maritime pine. The graphs represent the selection effect (fixation rate of non-synonymous mutations) vs. constraint effect (deleterious mutations).





Figure S3. Maximum likelihood three-species unrooted phylogenetic tree to estimate shared divergence between loblolly pine and Scots pine/maritime pine (see main text).

Example of branch length estimates using 168,534 bp of concatenated sequence per species (Pipn_1: one sequence for maritime pine, Pisy_1: one sequence for Scots pine and Pita_1: one sequence for loblolly pine).



Table S1. Sampling for the two conifer species. Numbers under candidate genes' column correspond to the exact number of samples per population for each of the six previously studied candidate genes. For CRSP and "extended CRSP" datasets the numbers show the maximum sample size per population.

Scots pine										
Site	Country	Latitude	Longitude	8 loc	8 loci (6 genes) from candidate genes					CRSP
				coL1	dhn1	dhn2	dhn5	gia	4cl	
Kolari	Finland	67.18	24.05	19	5	4	5	19	7	2
Usinsk	Russia	66.08	57.5	7				10	5	
Kaddekielas	Sweden	66.07	19.1	8				9	6	
Northern Sweden	Sweden	65.13	20.23							1
Northern Sweden	Sweden	64.15	16.07							1
Uusikaupunki	Finland	60.87	21.33	10	5	4	5	10	5	
Eastern Scotland	UK	57.05	-3.27	10				10	5	2
Kalsnava	Latvia	56.75	25.88	9				9	4	
Norra Gullabo	Sweden	56.47	15.92	10	5	5	5	10	6	
Radom	Poland	50.68	20.08	20	5	5	5	20	9	2
Haguenau	France	48.85	7.87	10	5	4	5	9	6	
Oberloisdorf	Austria	47.43	16.48		5	5	5			
Parma	Italy	44.62	10.15	10				9	3	2
Kalabak	Turkey	39.45	30.3		5	5	5			
Sierra de Baza	Spain	37.37	-2.83		5	4	5			2
Total (individuals)			-	113	40	36	40	115	56	12
Total (populations)				10	8	8	8	10	10	7

maritime pine											
Site	Country	Latitude	Longitude	gitude 8 loci (6 genes) from candidate genes				CRSP	CRIEC		
				coL1	dhn1	dhn2	dhn5	gia	4cl		
Landes	France	na	na							3	
Pleucadec	France	47.78	-2.34		5						
Mimizan (Landes)	France	44.13	-1.3		15	4	2		1		2
Galicia	Spain	na	na							1	
Unknown origin*	France	na	na							1	
Tabuyo del Monte	Spain	42.3	-6.22								2
San Cipriano de Ribarteme	Spain	42.12	-8.36		7						
Pinia (Corsica)	France	42.02	9.46	4	8	9	9	3	7		2
Pineta (Corsica)	France	41.96	9.04		6						
Coca	Spain	41.23	-4.5	13	17	15	13	10	16	2	4
Arenas de San Pedro	Spain	40.19	-5.12	12	19	15		10	13	1	4
Olba	Spain	40.17	-0.62	4		8	7	4	7	1	4
Quatretonda	Spain	38.97	-0.36	3		8	6	4	5	1	4
Cazorla	Spain	37.92	-2.92	9		7	7	9	9	1	4
Oria	Spain	37.52	-2.33	6	6	9	9	1	6	1	4
Tabarka	Tunisia	36.94	8.7	5	10	8	8	4	9	1	2
Cómpeta	Spain	36.85	-3.88			1	4		8		
Estepona	Spain	36.52	-5.12								2
Tamrabta	Morocco	33.6	-5.02	3	7	8	3	2	8	1	2
Sidi Meskour	Morocco	31.47	-6.83	1			1	2	3		
Total (individuals)			60	100	92	77	49	92	14	36	
Total (populations)				10	10	11	12	10	12	11	12

* Parent of a QTL mapping progeny with only approximate known origin (France)

Locus	Length (bp) ^a	Scots pine	maritime pine
4cl (exon 1, 2-3, 5) ^b	516+360+251 = 1127	this study ³ (accessions MF385276-MF385442)	Grivet et al. (2011)
dhn-1	636	Wachowiak et al. (2009) ^d	Eveno et al. (2008)
dhn-2	713	Wachowiak et al. (2009)	Grivet et al. (2011) ^c
dhn-5	559	Wachowiak et al. (2009)	Grivet et al. (2011)
coLl	3829	Pyhäjärvi et al. (2007); Kujala and Savolainen (2012) Pyhäjärvi et al. (2007); Kujala and Savolainen	this study ¹ (accessions MF385443-MF385502) this study ² (accessions
gia	1376	(2012)	MF385510-MF385558)

Table S2. Description of the six previously studied candidate genes (8 loci).

^aLength based on the common alignment between Scots pine and maritime pine, including indels.

^bExon_1, exon_2-3, exon_5 correspond to exon_c, exon_a, exon_b in Grivet et al. (2011).

^cThe candidate gene corresponds to *dhn2-Ps*.

^d*dhn-1* corresponds to *dhn-9* in Wachowiak et al. (2009).

¹ PCR reaction: for primer pairs Copr-promU2/ex1L1 and Copr-ex2U2/3utrL2 the 20 μl mix of reaction contained 0.8 mM dNTP, 0.5 μM of each primer, 1x Phusion HF Buffer (Phusion , Finnzymes), 25 ng DNA and 0.4 unit Taq polymerase (Phusion, Finnzymes); PCR conditions: 1 min at 98 °C, 35 cycles of 10 sec at 98 °C, 30 sec at 66 °C, 30 sec 72 °C, followed by 10 min at 72 °C (see primer's specification in Kujala and Savolainen 2012). For primer pairs Copr-109U/941L, Copr-718U/1919L and Copr1820U/1311L the 20 μl mix of reaction contained 0.5 μM of each primer, 1x Phusion Flash PCR Master Mix (Phusion Flash, Finnzymes) and 25 ng DNA. PCR conditions: 1 min at 98 °C, 35 cycles of 10 sec at 98 °C, 30 sec at 67 °C (Copr-109U/941L) or 63 °C (Copr-718U/1919L and Copr1820U/1311L), 30 sec at 72 °C, followed by 10 min at 72 °C (see primer's specification in Kujala and Savolainen 2012).

² PCR reaction: the 20 μl mix of reaction contained 0.8 mM dNTP, 0.5 μM of each primer, 1x Phusion HF Buffer (Phusion, Finnzymes), 25 ng DNA and 0.8 unit Taq polymerase (Phusion, Finnzymes). PCR conditions: 1 min at 98 °C, 35 cycles of 10 sec at 98 °C, 30 sec at 65 °C (ex11-U1/3utr-L1) or 66 °C (ex10-U1/ex11-L2), 1 min at 72 °C, followed by 10 min at 72 °C (see primer's specification in Kujala and Savolainen 2012). ³ PCR conditions are the same as in Kujala and Savolainen (2012), and primers are identical to those described in Grivet et al. (2011).

Locus	Function	E-value Pinus EST	E-value A. thaliana	Length (bp) ^a
<i>4cl (exon 1)</i> ^{<i>b</i>}	4-coumarate_CoA ligase (S)	0	7.00E-49	515
4cl (exon 2-3) ^b	4-coumarate_CoA ligase (S)	0	1.00E-24	359
4 <i>cl</i> (<i>exon</i> 5) ^{<i>b</i>}	4-coumarate_CoA ligase (S)	0	1.00E-19	249
dhn-1 ^c	Dehydrin (S)	0	5.00E-07	601
dhn-2 ^d	Dehydrin (S)	0	5.00E-03	573
dhn-5	Dehydrin (S)	0	6.00E-04	386
0_16976_02	3-ketoacyl-CoA synthase 6 (S)	na	1.00E-41	402
0_18745_02	Mitogen-activated protein kinase 4 (S)	2.00E-089	8.00E-17	685
0_2070_01	Heat stress transcription factor B-2b (S)	na	7.00E-11	411
0_3790_01	Phospholipase D alpha 1 (S)	3.00E-120	5.00E-35	512
0_4032_02	ARM repeat superfamily protein (S)	na	3.00E-28	390
0_4042_01	Glutathione S-transferase TAU 8 (S)	1.00E-169	2.00E-21	476
0_4285_01	Amino acid permease 3 (S)	0	3.00E-27	495
0_6683_01	Salt-inducible zinc finger 1 (S)	na	6.00E-10	459
0_6878_01	F-box leucine-rich repeat family protein MAX2 (S)	7.00E-047	5.00E-62	435
0_768_02	Putative protein kinase (S)	1.00E-109	1.00E-38	481
0_9082_01	Putative beta-1,3-endoglucanase (S)	0	2.00E-32	421
0_9524_02	U-box domain-containing protein 41 (S)	na	3.00E-25	449
0_990_01	Putative calcium-binding protein CML25 (S)	na	7.00E-30	399
2_1582_02	DNAJ heat shock protein-like protein (S)	0	9.00E-56	457
2_2931_01	Ethylene-responsive transcription factor RAP2.4 (S)	0	7.00E-33	454
2_3319_01	Autophagy-related protein 18D (S)	na	2.00E-13	325
2_3726_02	DNAJ heat shock protein-like protein (S)	na	1.00E-28	450

Table S3. 64 target genes related to stress responses (S) and phenology/photosynthesis (P).

2_6731_01	F-box protein GID2 (S)	na	4.00E-14	430
CL1524Contig1_03	Histidinol dehydrogenase (S)	5.00E -053	3.00E-17	439
CL1536Contig1_03	Mannose-1-phosphate guanylyltransferase (S)	5.00E -105	6.00E-37	207
CL2332Contig1_01	Calcium-dependent protein kinase 6 (S)	6.00E -057	3.00E-19	426
CL263Contig2_03	RNA-binding protein 47C' (S)	7.00E -047	6.00E-10	444
CL3771Contig1_04	Ubiquitin-conjugating enzyme E2 32 (S)	9.00E -091	3.00E-11	446
UMN_2399_01	U-box domain-containing protein 13 (S)	0	8.00E-27	437
UMN_5272_01	6-phosphogluconate dehydrogenase, decarboxylating 3 (S)	na	3.00E-83	444
UMN_CL132Contig1_03	Malate dehydrogenase (S)	7.00E -026	3.00E-62	319
2_9480_01	Malate dehydrogenase (S)	na	2.00E-63	420
2_1014_01	Heat stress transcription factor B-1 (S)	1.00E -059	2.00E-06	431
0_1123_01	Heat shock protein 70B (S)	0	1.00E-38	404
0_11591_01	Protein auxin RESPONSE 4 (S)	0	5.00E-33	419
0_11649_01	Tubulin beta-8 chain (S)	6.00E -128	4.00E-84	572
0_11649_03	Tubulin beta-2/beta-3 chain (S)	0	2.00E-72	344
0_11684_01	Coronatine-insensitive protein 1 (S)	0	7.00E-37	491
0_12117_01	Adenine nucleotide alpha hydrolase-like protein (S)	0	2.00E-15	414
0_12896_01	F-box protein SKIP2 (S)	2.00E -136	5.00E-55	434
0_143_01	Peroxidase 15 (S)	na	8.00E-22	434
0_17010_02	Putative UDP-glucose 6-dehydrogenase 1 (S)	0	1.00E-76	375
0_4588_01	Aldehyde dehydrogenase 2B4 (S)	na	1.00E-14	218
CL1029Contig1_01	Putative galactinolsucrose galactosyltransferase 2 (S)	3.00E -085	3.00E-13	427
CL305Contig1_05	Dihydrolipoyl dehydrogenase 1 (S)	na	6.00E-44	302
CL3795Contig1_01	Amino acid dehydrogenase family protein (S)	na	1.00E-21	545
0_7921_01	Glucose and ribitol dehydrogenase homolog 1 (S)	na	2.00E-17	336
0_15991_01	E3 ubiquitin-protein ligase COP1 (P)	na	3.00E-17	317
coLl	Zinc finger protein Constans-like 3 (P)	0	8,00E-35	3797
gia	Gigantea protein (P)	0	1,00E-27	1287
0_12156_01	Inactive leucine-rich repeat receptor-like protein kinase	0	4.00E-47	431

	CORYNE (P)			
0_12156_02	inactive leucine-rich repeat receptor-like protein kinase CORYNE (P)	0	7.00E-37	450
0_16400_01	Protein UNUSUAL FLORAL ORGANS (P)	0	1.00E-49	437
UMN_3408_01	histone-binding protein RBBP4 (P)	0	2.00E-26	433
0_3723_01	STRUBBELIG-receptor family 3 (P)	2.00E -072	3.00E-16	572
0_7454_01	Probable serine/threonine-protein kinase (P)	na	9.00E-17	454
0_8850_02	Photosystem I P700 chlorophyll a apoprotein (P)	na	1.00E-52	320
2_6995_01	Phosphoenolpyruvate carboxylase 4 (P)	0	3.00E-61	407
UMN_3561_02	Photosystem II 47 kDa protein (P)	0	3.00E-41	352
UMN_5101_03	Cytochrome b6/f complex subunit V (P)	na	2.00E-13	402
UMN_6852_02	Cytochrome f (P)	0	1.00E-37	397
UMN_6924_03	Photosystem II 47 kDa protein (P)	na	2.00E-42	321
CL1430Contig1_06	Pyrophosphatefructose 6-phosphate 1-phosphotransferase subunit alpha 1 (P)	1.00E -056	4.00E-19	619

^aLength based on the common alignment between Scots pine and maritime pine, including indels. ^b*exon-1*, *exon-2-3*, *exon-5* correspond to *exon-c*, *exon-a*, *exon-b* respectively in Grivet et al. (2011). ^c*dhn-1* corresponds to *dhn-9* in Wachowiak et al. (2009). ^d*dhn-2* corresponds to *dhn2-Ps* in Grivet et al. (2011).

Compound reference locus	Individual reference locus
LG1_a	0_18018_01, CL3054Contig1_01, UMN_1023_01
LG1_b	0_15329_01, 0_17206_01, 0_17607_02
LG2_a	2_9455_01, 0_8531_01, 0_9091_01, UMN_5867_01
LG2_b	0_13929_02, 0_16732_01, 0_2217_01, 2_9087_01
LG3_a	0_18261_01, 0_7001_01, UMN_3006_01, UMN_3444_01, UMN_927_01
LG3_b	0_17082_01, 0_5575_01, 0_9448_01, 2_1528_01
LG3_c	0_4756_01, 2_6618_01, 0_11270_01, 0_12683_01, 0_846_01, CL814Contig1_06
LG4_a	0_16068_01, 0_9444_01, CL4511Contig1_02
LG4_b	0_13383_01, 0_7171_01, 2_3591_03, CL1045Contig1_01, CL1238Contig1_01, UMN_1037_01
LG5_a	0_10453_01, 0_6465_01, CL3037Contig1_06, CL3758Contig1_05, 0_18350_01
LG5_b	0_10054_01, 2_3941_01, 2_5724_02, 2_9603_01, CL415Contig1_04
LG5_c	UMN_4904_01, 0_236_01, 2_2936_01, CL4342Contig1_01, UMN_801_01
LG6_a	0_9383_01, 2_7725_01, 2_8852_01, CL1004Contig1_08, CL4432Contig1_04
LG6_b	0_11980_01, 0_12929_02, 0_8359_01, 0_8844_01, 0_9329_02, 2_5064_01, CL544Contig1_03
LG7_a	0_10667_02, 2_5996_01, CL1848Contig1_01, CL572Contig1_02
LG7_b	0_1659_02, 0_18470_01, 0_2078_01, 2_5636_01, CL4470Contig1_01
LG7_c	0_14976_01, 0_4105_01, 0_4394_01, 2_6491_01, 2_9291_02
LG8_a	0_10267_01, 0_14221_01, 2_3947_01, CL1455Contig1_07
LG8_b	0_17127_01, 0_6999_01, 2_2960_02, CL1698Contig1_01, CL3539Contig1_01
LG9_a	0_17143_02, 2_10236_01, CL1694Contig1_04, UMN_6426_02
LG9_b	2_5099_01, 2_7852_01, 2_9930_01
LG9_c	0_13841_01, 0_16459_01, 2_684_01, 2_974_01
LG10_a	0_12021_01, 0_12978_02, 2_4724_01, 2_6130_01
LG10_b	0_13484_01, 0_16860_01, 2_6052_01, UMN_5833_01
LG11_a	0_12190_02, 0_16009_01, 0_17247_02, 2_7918_01, CL2472Contig1_01
LG11_b	0_16889_02, 0_2433_01, 0_5204_01, CL4023Contig1_01

 Table S4. 28 compound reference loci obtained using linkage map information.

LG12_a	CL905Contig2_01, 0_16169_01, 0_2885_01, 0_3261_01
LG12_b	0_9922_01, 0_11090_01, 1_5675_01, 2_10212_01, UMN_2174_01

Locus	Scots	pine		maritim	e pine
-	Empirical		_	Empirical	
	Bayes	Bayesian		Bayes	Bayesian
	SnIPRE	SnIPRE	_	SnIPRE	SnIPRE
dhn1 (S)	negative	negative			positive
<i>dhn2</i> (S)	positive	positive			
dhn5 (S)	positive	positive			
coL1 (P)	positive	positive		positive	positive
4cl_exon1 (S)	positive				
0_4042_01 (S)				positive	positive
2_9480_01 (S)				negative	negative
0_143_01		positive			
0_3723_01	negative				
0_4032_02				positive	
0_4588_01		positive			
0_9082_01	positive				
0_11684_01		positive			
0_11649_01				negative	
CL1430Contig1_06		positive			
Total positive	5	7		3	3
Total negative	3	2		2	1
Total	8	9		5	4

Table S5. Loci found under selection with the Empirical Bayes SnIPRE and the Bayesian SnIPRE methods (for silent sites).

Table S6. Count of fixed differences between loblolly pine and Scots pine/maritime pine. "Originated in one species" refers to mutations that arose only in Scots pine or in maritime pine.

	Scots pine				maritime pir	ne
	All sites	0-fold sites	4-fold sites	All sites	0-fold sites	4-fold sites
Fixed sites	1868	402	333	2199	446	412
Shared sites	1239	250	227	1239	250	227
Originated in one species	629	152	106	960	196	185

Table S7. Bootstrapping procedure to compare nucleotide diversity estimates in loci transferred from loblolly pine to European pines (364 CRSP loci, N=14) with those randomly resampled from loblolly pine genome (about 6,000 loci available). Numbers in bold indicate significant departure from random set of loci.

	S total	π total	$\theta_{_W}$ total	π syn	$\theta_{_W}$ syn	π nsyn	θ_w nsyn
Transferred loci set	3.36	0.00213823	0.00253591	0.00480557	0.00560486	0.00093046	0.00112435
Low 95CI	3.02	0.00188983	0.00225759	0.00412203	0.00490509	0.00074009	0.00091097
High 95CI	3.70	0.00238664	0.00281423	0.00548911	0.00630463	0.00112082	0.00133772
1,000 bootstraps (364 loci)							
Low 95CI	3.38	0.00307339	0.00342669	0.00466651	0.00524553	0.00103738	0.00122625
High 95CI	4.48	0.00439749	0.00481324	0.00800163	0.00884918	0.00232990	0.00244981

S total: number of segregating sites per locus; π (total, syn, nsyn): Tajima's nucleotide diversity (Tajima 1989) for total sites, synonymous sites, and non-synonymous sites; θ_w (total, syn, nsyn): Watterson's nucleotide diversity (Watterson 1975) for total sites, synonymous sites, and non-synonymous sites; 95Cl: 95% Confidence Intervals.

Table S8. Summary statistics for all sites, synonymous sites, non-synonymous sites and silent sites, for the two conifer species for all loci (target and reference genes) for the CRSP and for the "extended CRSP" datasets. Divergence estimates are given using loblolly pine as reference. Statistics were normalized for varying sample size across loci with the SDMTools package in R.

	Scots pine (CRSP)					
	all	syn	nsyn	silent		
θ_w (stdev)	0.003302 (0.003664)	0.006871 (0.008962)	0.001470 (0.002685)	0.005588 (0.006476)		
π (stdev)	0.002926 (0.003886)	0.006346 (0.009822)	0.001254 (0.002494)	0.004978 (0.006712)		
Divergence (JC) (stdev)	0.016598 (0.010255)	0.032276 (0.025316)	0.006973 (0.008064)	0.028613 (0.018187)		

	maritime pine (CRSP)					
	all	syn	nsyn	silent		
θ_w (stdev) π (stdev) Divergence (JC) (stdev)	0.002141 (0.002162) 0.002236 (0.002507) 0.019092 (0.014916)	0.004189 (0.005440) 0.004426 (0.006942) 0.037465 (0.031322)	0.000823 (0.001393) 0.000905 (0.001803) 0.007029 (0.007974)	0.003749 (0.003783) 0.003785 (0.004145) 0.033709 (0.025498)		

	maritime pine ("extended CRSP")						
	all	syn	nsyn	silent			
$ \theta_{w} (stdev) $ $ \pi (stdev) $ Divergence (JC) (stdev)	0.002588 (0.001915) 0.002912 (0.00252) 0.021141 (0.015417)	0.00463 (0.005147) 0.00543 (0.006992) 0.041243 (0.036966)	0.001016 (0.00138) 0.001083 (0.001996) 0.008142 (0.009089)	0.004193 (0.003292) 0.004768 (0.00428) 0.037145 (0.030381)			

Table S9. Nucleotide diversity for maritime pine with CRSP (364 loci + 8 previously studied loci = 372 loci) and "extended CRSP" (128 + 8 previously studied loci = 136 loci) datasets using bootstrap resampling with replacement for different number of loci (indicated in brackets).

	CRSP		"extended CRSP"	
	π silent	θ_w silent	π silent	θ_w silent
-				
Full dataset	0.003785	0.003749	0.004768	0.004193
Lower 95% CI	0.003282	0.003174	0.004076	0.003580
Higher 95% CI	0.004337	0.004358	0.005530	0.004770
Mean 1,000 bootstrats (10)				
Lower 95% CI	0.001220	0.001436	0.002448	0.002319
Higher 95% CI	0.007082	0.007389	0.007804	0.006756
Mean 1,000 bootstrats (50)				
Lower 95% CI	0.002380	0.002398	0.003612	0.003248
Higher 95% CI	0.005065	0.005451	0.005902	0.005192
Mean 1,000 bootstrats (100)				
Lower 95% CI	0.002783	0.002826	0.003939	0.003579
Higher 95% CI	0.004773	0.004946	0.005631	0.004904

Statistics were normalized for varying sample size across loci with the SDMTools package in R.